VALID INFERENCE FOR GROUP-ANALYSIS OF FUNCTIONALLY ALIGNED FMRI IMAGES

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CON40 - IMAGE ANALYSIS AND VISUAL METHODS



Multi-subjects fMRI studies permit to compare studies across subjects, to generalize and to validate the results.

The anatomical and functional structure of brains vary across subjects even in response to identical sensory inputs.



Anatomical Alignment (MNI normalization) ¹;
Functional Alignment.

¹Jenkinson, M. and Smith, S., Medical image analysis (2001)



We can assume that the neural activities in different brains are **noisy rotations of a common space**.

Analyzing group activation means performing roughly 200, 000 statistical tests equal to the number of voxels \rightarrow **multiple testing problem.**

We propose a method that simultaneously

- resolves the functional misalignment;
- gains power in group analysis inference without affecting the type I error.

The functional alignment is done by the **ProMises**² (von Mises-Fisher-Procrustes) model

sorts the null hypotheses based on a priori information independent from the test statistics \rightarrow valid inference.

²Andreella, A. and Finos, L., Psychometrika (2022).

PROMISES MODEL

Let *n* time points, *m* voxels and i = 1, ..., N subjects:

$$\mathbf{X}_i = \alpha_i (\mathbf{M} + \mathbf{E}_i) \mathbf{R}_i^{\mathsf{T}}$$

where

- $\{\mathbf{X}_i \in \mathbb{R}^{n \times m}\}_{i=1,...,N}$ represent the matrices to be aligned;
- **M** $\in \mathbb{R}^{n \times m}$ is the configuration reference matrix;

$$\blacksquare \vec{E}_i \sim \mathcal{N}_{nm}(\vec{O}, \Sigma_m \otimes \Sigma_n);$$

■ { $\mathbf{R}_i \in \mathcal{O}(m)$ }_{*i*=1,...,*N*} orthogonal matrix parameter, and $\alpha_i \in \mathbb{R}_{>0}$ scaling parameter.

R_i ~ **von Mises-Fisher**³ distribution with location parameter

 $\mathbf{F} \in \mathbb{R}^{m \times m}$ and concentration parameter $k \in \mathbb{R}_{>0}$, i.e.,

$$f(\mathbf{R}_i) = C \exp(\frac{\mathbf{k}}{\mathbf{r}} \operatorname{tr}(\mathbf{F}^{\mathsf{T}} \mathbf{R}_i))$$

³Downs, T. D., Biometrika (1972).

Theorem

Let the singular value decomposition of $\mathbf{X}_{i}^{\mathsf{T}} \mathbf{\Sigma}_{n}^{-1} \mathbf{M} \mathbf{\Sigma}_{m}^{-1} + k\mathbf{F}$ be $\mathbf{U}_{i} \mathbf{D}_{i} \mathbf{V}_{i}^{\mathsf{T}}$, then the ProMises model returns: 1. $\hat{\mathbf{R}}_{i}$ equals $\mathbf{U}_{i} \mathbf{V}_{i}^{\mathsf{T}}$; 2. $\hat{\alpha}_{i \hat{\mathbf{R}}_{i}}$ equals $\frac{\|\mathbf{X}_{i} \hat{\mathbf{R}}_{i}\|_{F}^{2}}{\mathsf{Tr}(\mathbf{D}_{i})}$.

The von Mises-Fisher distribution is conjugate⁴ to the matrix normal distribution, with location parameter:

$$\mathbf{X}_{\mathbf{i}}^{\mathsf{T}} \mathbf{\Sigma}_{\mathbf{n}}^{-1} \mathbf{M} \mathbf{\Sigma}_{\mathbf{m}}^{-1} + k \mathbf{F}.$$

⁴Green, P. J. and Mardia, K. V., Biometrika (2006).

Let consider a given voxel k, where k = 1, ..., m for each subject i:

$$\boldsymbol{X}_i = \boldsymbol{D}\boldsymbol{B}_i + \boldsymbol{Z}\boldsymbol{G}_i + \boldsymbol{E}_i$$

where

- **D** $\in \mathbb{R}^{n \times p}$ and **Z** $\in \mathbb{R}^{n \times q}$ are fixed matrices;
- **B**_i ∈ ℝ^{p×m}, **G**_i ∈ ℝ^{q×m} and **B**_i = **B** + **U**_i, **G**_i = **G** + **g**_i with **B** is the true matrix of **fixed effects of interest**, **G** of fixed nuisance effects;
- $\blacksquare [\boldsymbol{U}_{i}^{\mathsf{T}} | \boldsymbol{g}_{i}^{\mathsf{T}}] \sim \mathcal{M}\mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma}_{m}, \boldsymbol{\Sigma}_{pq}) \text{ is the matrix of random effects.}$

We have now a set of N matrices $\{\hat{B}_1, \ldots, \hat{B}_N; \hat{B}_i \in \mathbb{R}^{p \times m}\}$ (e.g., describing the difference between the neural activation during two stimuli recorded in the voxel $k \in \{1, \ldots, m\}$ of the subject *i*).

One-sample t-test

$$T = \frac{\hat{\mu}}{\hat{\sigma}/\sqrt{N}}$$
(1)

• $\hat{\mu} = \sum_{i=1}^{N} \hat{B}_i / N$ is the sample mean between-subjects with $\hat{\mu} \in \mathbb{R}^{p \times m}$;

• $\hat{\sigma} = \frac{1}{N-1} \sum_{i=1}^{N} (\hat{B}_i - \hat{\mu})^2$ is the sample variance between-subjects with $\hat{\sigma} \in \mathbb{R}^{p \times m}$.

One local test T_k for each $H_0^k : \mu_k = 0$ vs the two-sided alternative hypothesis \rightarrow one T_k for each k voxel.

Theorem

Let consider the p-values p_k related to the statistical test T_k , where $k \in M = \{1, ..., m\}$. If the ProMises model is valid, then:

$$\Pr(p_{k} \leq \alpha \mid \hat{R}_{i}, \hat{\alpha}_{i}) = \alpha, \quad \forall k \in S$$

where S \subseteq M is the set of true null hypotheses, and α is the significance level.

The information involved in the estimation of \mathbf{R}_i and \mathbf{B}_i are orthogonal under H_0 .

- We align the images of of 18 subjects passively listening to vocal, i.e., speech, and non-vocal sounds.
- After the X_i matrices' alignment, the one-sample t-test was performed to study the significance of the group's mean activation concerning the difference between the neural activation during the two stimuli.
- The inferential analysis is performed on the superior temporal gyrus. The ProMises model is compared with raw data (i.e., anatomical alignment only).



Using data functionally aligned, we have tests **85.85%** higher than those returned by using raw data, with baseline 50%.

- We proved the validity of the inference in functionally aligned data;
- Alignment reduces the intra-subject variability and localizes the signal → more powerful inference;
- One sample t-tests at the group level using data aligned by the ProMises model illustrate higher absolute values than the one-sample t-tests computed using raw data.

The method is implemented as

- Python module: ProMisesModel;
- R package: alignProMises

on my GitHub profile https://github.com/angeella.